



# Quantitative economic geography meets history: Questions, answers and challenges<sup>☆</sup>



Dávid Krisztián Nagy

CREi, Universitat Pompeu Fabra, Barcelona GSE, Spain

## ARTICLE INFO

### Keywords:

Economic geography  
Regional economics  
Urban economics  
Economic history

## ABSTRACT

A rapidly growing literature uses quantitative general equilibrium models of economic geography to study the economic impact of historical events such as the railroad revolution, industrial take-off, structural transformation and wars. I identify three key challenges facing this literature: the tractability of model structure, the availability of historical data, and issues related to identification. I review the literature by discussing how it has been addressing each of these challenges. While doing so, I point out the rich set of questions that this literature can address, as well as the methodological innovations it has conducted to answer these questions.

## 1. Introduction

How much economic growth was induced by railroads? How much did the falling costs of trade contribute to the current world income distribution? What was the economic impact of city bombing, of war-induced border changes, or population expulsions? How did 19th-century pollution reshape developed-world cities? Each of these questions calls for studying the economic consequences of large-scale historical events. And, although these events took place in the past, they might be relevant for today's economists for at least two reasons. First, they might provide explanations for why the current economy looks how it looks. For instance, understanding the origins of the world income distribution has been a long-held goal among researchers (Acemoglu, 2009). Second, similar large-scale events may take place in the future, and economists and policymakers naturally want to understand what they will induce. Pollution, for instance, is at least as much of a problem for Chinese cities today as for 19th-century cities in the developed world (Ebenstein et al., 2015).

What is also common across these large-scale historical events is that they are *inherently geographic*. Railroads connect certain locations, but not others. Bombing and pollution affect certain parts of a city but might leave other parts intact. Border changes or population expulsions

happen across certain points in space. This is not to say that other locations in the economy are unaffected by these events. They might experience indirect effects through their *spatial linkages* – trade, commuting, or factor mobility – with directly affected locations. As a result, any methodology aimed at studying the effects of these geographic events needs to embrace the notion of distinct locations connected through spatial linkages. Traditional general equilibrium macroeconomic models do not have this feature, as they model the economy as a dimensionless point in space. This is precisely why researchers interested in inherently geographic questions moved towards modeling the economy as a set of locations, linked to one another through trade and factor mobility (Fujita and Thisse, 2002).

Another prominent methodology used to study the effects of geographic events is reduced-form empirics. A clear advantage of this approach over structural modeling is that it is devoid of the issue of model misspecification, which lends additional credibility to its results.<sup>1</sup> At the same time, the core principle of this methodology is the comparison of individuals or locations affected by an event (the “treated group”) to those unaffected by the event (the “control group”). In a world with spatial linkages, it is likely that a large-scale event affects every location. In this case, reduced-form empirics can identify the effect of the event on locations designated as “treated” *relative to*

<sup>☆</sup> I am grateful to Stephan Heblich, Giacomo Ponzetto, Stephen Redding and two anonymous referees for helpful comments and suggestions. I acknowledge financial support from the Spanish Ministry of Economy and Competitiveness, through the Severo Ochoa Programme for Centres of Excellence in R&D (CEX2019-000915-S), through the Juan de la Cierva Grant (FJCI-2017-34728), and from the Government of Catalonia, through CERCA and SGR Programme (2017-SGR-1393). All errors are my own.

E-mail address: [dnagy@crei.cat](mailto:dnagy@crei.cat).

<sup>1</sup> This argument is made, for instance, in Baum-Snow and Ferreira (2015), who also discuss the methodological challenges faced by the empirical economic geography literature.

<https://doi.org/10.1016/j.regsciurbeco.2021.103675>

Received 1 November 2020; Received in revised form 1 March 2021; Accepted 23 March 2021

Available online 5 April 2021

0166-0462/© 2021 The Author(s). Published by Elsevier B.V. This is an open access article under the CC BY-NC-ND license (<http://creativecommons.org/licenses/by-nc-nd/4.0/>).

locations designated as “control.” As described in Redding and Turner (2015), such a relative effect might reflect the *relocation* of existing economic activity between treated and control locations, rather than actual (positive or negative) *growth* as a result of the event. Reduced-form empirical techniques cannot distinguish between such growth and relocation effects. As a consequence, unlike structural general equilibrium models, they cannot be used to measure the effects on the economy as a whole.<sup>2</sup>

Quantitative general equilibrium models of economic geography are able to combine the advantages of reduced-form empirics and structural macroeconomic models. On the one hand, they embrace the realism of multiple locations connected through complex spatial linkages, unlike models that are dimensionless or feature stylized geography. This allows the researcher to combine the model with rich spatial data, or even with reduced-form empirical strategies to identify certain moments from the data, as I also illustrate in the quantitative example of Section 2.2. On the other hand, quantitative modeling, structural by nature, allows the researcher to measure the effects of large-scale events on the whole economy and distinguish growth from the mere relocation of economic activity.

Needless to say, these advantages of quantitative modeling come at a cost. In particular, using quantitative models to study history imposes three key challenges on the researcher. One of these challenges is that the realism of model structure often comes at the expense of tractability. Another stems from the sparse nature of historical data on macroeconomic aggregates. A final challenge, shared with reduced-form empirics, is that identifying the effects of events may be subject to biases due to omitted variables and endogeneity.

In the remainder of this article, I review the rapidly growing literature that uses quantitative general equilibrium models of economic geography to study the economic impact of large-scale historical events. I do so by discussing the above three challenges and the ways the literature has been addressing each of them. Thus, Section 2 discusses the challenge of model tractability, Section 3 focuses on the scarcity of historical data, while Section 4 looks at issues related to identification. Section 5 concludes the article by suggesting avenues for further research in the field. This paper is closely related and complementary to Hanlon and Heblich (2020), who review the new insights brought by studies in the intersection of economic geography and history into classical questions in urban economics. The paper is also related by Redding and Rossi-Hansberg (2017), who offer a general overview of quantitative economic geography models.

## 2. Model tractability

Incorporating the complex nature of real-world spatial interactions often comes at the expense of tractability in quantitative modeling. To highlight which model ingredients do and which do not prove challenging from a tractability perspective, I proceed in this section as follows. In Section 2.1, I present a class of quantitative spatial models with trade, labor mobility and commuting that, even though not solvable in closed form, prove tractable in various respects. Specifically, (1) the existence and uniqueness of the model’s equilibrium can be characterized theoretically; (2) a simple algorithm can be used to solve for the equilibrium on the computer; (3) unobserved location fundamentals that rationalize the observed data as an equilibrium can be uniquely recovered for given structural parameters (that is, the model can be *inverted*). In Section 2.2, I develop a quantitative application to illustrate that this class of

<sup>2</sup> A related issue is that, by focusing on direct treatment effects, reduced-form empirical studies often overlook indirect effects transmitted through spatial linkages across locations. This issue, in principle, can be solved in the reduced form, either by investigating spillover effects to neighboring locations or by redefining treatment variables to capture both direct and indirect effects (Donaldson and Hornbeck, 2016).

models is amenable to answering important historical questions. This application showcases the strengths of these models when it comes to their practical use, but also their limitations. Finally, in Section 2.3, I consider additional model ingredients absent from the class of models presented in Section 2.1: dynamics, multiple sectors, and endogenous infrastructure development. I discuss how these additional model ingredients could increase model realism in the context of the application of Section 2.2. At the same time, I highlight how the additional ingredients pose challenges to tractability. I also point out how existing quantitative historical studies that incorporate some of these additional ingredients have addressed these challenges.

### 2.1. A tractable class of quantitative spatial models

In this section, I develop a quantitative spatial model of a set of locations, where locations are linked through trade, labor mobility and commuting (Section 2.1.1).<sup>3</sup> Next, I show that this model features a tractable structure; more precisely, the existence and uniqueness of the model’s equilibrium can be characterized theoretically, the equilibrium can be computed using a simple algorithm, and the model can be inverted to recover unobserved location fundamentals that rationalize the observed data as an equilibrium (Section 2.1.2). Finally, I show isomorphisms between the model and a set of other models, some of which have been used in the quantitative economic geography literature (Section 2.1.3). As a result of these isomorphisms, the tractability of model structure I show for the model of Section 2.1.1 applies to this entire class of quantitative spatial models.

#### 2.1.1. Model setup

The economy consists of a discrete set of  $S$  locations, indexed by  $r$ ,  $s$ , or  $u$ .  $\bar{L} > 0$  workers inhabit the economy, where  $\bar{L}$  is exogenously given. There is one sector producing tradable goods. Within this sector, each location produces one good that workers view as different from the goods produced at other locations.<sup>4</sup> Hence, I index each tradable good by the index of its production location. Besides tradables, workers also consume housing, a homogenous nontradable good that is available in exogenous positive supply at each location. Housing payments go to immobile local landlords, who spend their entire income on tradables and have the same preferences over tradables as workers (Monte et al., 2018).

**Consumption and location choice.** Each worker chooses a residential location to live. They also choose a – potentially different – workplace location, where they inelastically supply the one unit of labor they own. Workers are atomistic, implying that they take wages, the prices of tradables and the price of housing as given at every location.

If worker  $i$  chooses to live at location  $r$  and work at location  $s$ , she obtains utility

$$U_i(r, s) = a_i(r, s) \kappa(r, s)^{-1} \left[ \sum_{u=1}^S q_i(r, s; u)^{\frac{\sigma-1}{\sigma}} \right]^{\frac{\sigma}{\sigma-1}} H_i(r, s)^{1-\nu}$$

where  $a_i(r, s)$  denotes the amenities enjoyed by the worker at  $r$  and  $s$ ,  $\kappa(r, s) \geq 1$  is the cost of commuting between residence  $r$  to workplace  $s$ ,<sup>5</sup>  $H_i(r, s)$  is the quantity of housing consumed by the worker at her residence  $r$ , and  $q_i(r, s; u)$  is her consumption of tradable good  $u$ , where  $\sigma$  is the elasticity of substitution across goods. I make the

<sup>3</sup> The model is closest in its structure to Monte et al. (2018), Heblich, Redding and Sturm (2020), and Allen and Arkolakis (2014).

<sup>4</sup> In trade and geography, this is called the Armington assumption (Anderson, 1979). Section 2.1.3 shows isomorphisms between the model and alternative models in which I relax this assumption.

<sup>5</sup> Commuting between two locations can be infinitely costly, in which case  $\kappa(r, s) = \infty$ . Thus, the model embeds no commuting as a special case ( $\kappa(r, s) = \infty$  for any  $r \neq s$  and  $\kappa(r, r) = 1$ ).

standard assumption that  $\sigma > 1$ . That is, tradables are imperfect substitutes.

Amenities are drawn from a Fréchet distribution that is independent across workers and residence-workplace pairs. This is a standard assumption in the literature (Ahlfeldt et al., 2015; Redding, 2016), and captures the idea that workers have different idiosyncratic tastes for different locations. More precisely, I assume that the cumulative distribution function of  $a_i(r, s)$  takes the form

$$Pr [a_i(r, s) \leq a] = e^{-\bar{a}(r,s)N(r)^{-\lambda}a^{-1/\eta}}$$

where  $\bar{a}(r, s) > 0$  is the exogenous fundamental amenity level of the pair  $(r, s)$ ,  $N(r)^{-\lambda}$  is a congestion disamenity that depends on the population of the residential location  $r$ ,  $N(r)$ , with elasticity  $\lambda \geq 0$ ,<sup>6</sup> and  $\eta \in (0, 1)$  is a parameter driving the dispersion in idiosyncratic location tastes.

**Production and trade.** Each tradable good is produced by a large number of perfectly competitive firms that face a constant returns to scale production technology. As a result, there exists a representative firm. The representative firm of location  $s$  faces the production technology

$$q(s) = A(s)L(s)$$

where  $q(s)$  is the firm's output,  $A(s)$  is the location's productivity that the firm takes as given, and  $L(s)$  is employment. Employment at  $s$  potentially has an effect on productivity:

$$A(s) = \bar{A}(s)L(s)^\alpha,$$

a relationship that the firm does not internalize. This is a formulation of agglomeration externalities that is standard in the literature (Ciccone and Hall, 1996; Allen and Arkolakis, 2014).  $\bar{A}(s) > 0$  is the fundamental productivity level of location  $s$ , which is exogenous.

Trade in tradable goods is subject to *iceberg* costs. That is, if tradable goods start to be shipped from a location to another, only a fraction of these goods arrives. In particular,  $\tau(s, r) \geq 1$  units of good  $s$  need to be shipped from the production location  $s$  so that one unit arrives at  $r$ . The remaining units melt away in transit. As firms take all prices as given, no arbitrage guarantees that the ratio of the good's price between  $r$  and  $s$  also equals  $\tau(s, r)$ , as long as the good is traded between these two locations.<sup>7</sup>

**Equilibrium.** An equilibrium of the model is a set of prices and quantities such that workers choose their consumption levels as well as residence and workplace locations to maximize their utility; landlords choose their consumption levels to maximize their utility; firms choose their employment and output to maximize their profits; and labor, housing and tradable goods markets clear. Section A.1 of the Online Appendix shows that the equilibrium conditions of the model can be reduced to a system of  $5S + S^2 + 1$  equations. To compute the equilibrium, the researcher needs to solve these equations for  $5S + S^2 + 1$  endogenous variables; namely, for the residential population, the employment level, the wage level, the spending on tradables and the ideal price index at each location, the commuting flows between each pair of locations, as well as the economy-wide level of workers' expected utility.

### 2.1.2. Model tractability: theoretical results

This section presents three theoretical results that highlight the tractability of the model. Theorem 1 shows that the existence and uniqueness of the model's equilibrium are guaranteed by a condition that only depends on the model's structural parameters. As I argue,

<sup>6</sup> A generalization of the model would allow for the disamenity to depend on both residential and workplace population:  $N(r)^{-\lambda_1}L(s)^{-\lambda_2}$ , where  $L(s)$  denotes the number of people working at  $s$ . All the theoretical results of Section 2.1.2 carry over to this more general case.

<sup>7</sup> Similar to commuting costs, trade can be prohibitively costly across certain pairs of locations, in which case  $\tau(s, r) = \infty$ .

this condition tends to hold if the model's agglomeration force is not too strong. Next, Theorem 2 shows that the theoretical condition for equilibrium uniqueness also guarantees that a simple algorithm can be used to solve for the model's equilibrium on the computer. Finally, Theorem 3 shows that, for a given set of structural parameters, the model inversion identifies a unique set of unobserved amenities and productivities (up to scale) that rationalize the observed data as an equilibrium.

In the proofs of Theorems 1 to 3, I rely heavily on the set of theoretical results presented in Allen et al. (2020). For brevity, the proofs are relegated to Section A.2 of the Online Appendix.

**Theorem 1.** *The model's equilibrium exists and is unique under a condition that only depends on the model's structural parameters.*

The Online Appendix presents the specific condition under which equilibrium existence and uniqueness are guaranteed. I have not included the condition in the main text as it is a complex function of matrices whose entries depend on the model's structural parameters. Intuitively, the condition is more likely to hold if the model's agglomeration force (the agglomeration externality in production) is weak relative to the congestion forces (housing, the congestion disamenity, and the dispersion of idiosyncratic location tastes). This result should not be surprising. Under strong agglomeration forces, the concentration of economic activity at a certain location can sustain itself, which may give rise to different equilibria in which concentration arises at different locations.

If the equilibrium is guaranteed to be unique, then the researcher can be sure that a numerical procedure that has found an equilibrium has found the *only* equilibrium of the model. In general, however, a uniqueness result does not suggest a numerical procedure that can be used to find the equilibrium. That said, the following theorem suggests such a procedure. Moreover, it shows that the procedure is guaranteed to find the equilibrium if the condition for uniqueness in Theorem 1 holds.

**Theorem 2.** *Assume that the sufficient condition for existence and uniqueness from Theorem 1 holds. Then a simple iterative algorithm is guaranteed to converge to the equilibrium spatial distribution of residential population, employment, wages, spending on tradables and price indices. Finally, workers' expected utility and commuting flows can be obtained in closed form as a function of the former distributions.*

The simple iterative algorithm, as shown in Section A.2 of the Online Appendix, consists of guessing any initial distribution of residential population, employment, wages, spending on tradables and price indices, and then updating these distributions using the model's equilibrium conditions until convergence.

Although the convergence of the simple iterative algorithm is guaranteed by Theorem 2, the algorithm would be of little practical use if it were slow. However, Allen et al. (2020) show that the rate of convergence for such algorithms depends on how far the values of structural parameters are from the boundary of the region in which the condition of Theorems 1 and 2 holds. In practice, such algorithms tend to be very quick unless the values of structural parameters are extremely close to this boundary.

Theorems 1 and 2 provide a powerful tool to the researcher. As long as the values of the model's structural parameters guarantee uniqueness (Theorem 1), the researcher can use the simple iterative algorithm of Theorem 2 to compute the equilibrium spatial distribution of economic activity under any trade costs, commuting costs, and so on. For instance, she can simulate a counterfactual world in which a new railroad is built to connect particular locations.<sup>8</sup>

<sup>8</sup> This being said, the fact that the model features a unique equilibrium is not necessarily appealing in every context. In particular, the economic geography literature has argued for the indeterminacy of spatial structure in certain contexts. For details, see the review article by Lin and Rauch (2020).

Of course, the results of such counterfactual exercises depend on the underlying distribution of fundamental amenities and productivity. In practice, these fundamentals are unobserved. Thus, it is important to recover their spatial distributions before conducting counterfactual exercises with the model. A standard method for this consists of finding the distributions of fundamentals that rationalize the real-world data as an equilibrium (*model inversion*).

To invert the model, the researcher should ideally know if there exists a unique set of fundamentals that rationalize the observed data. Under uniqueness, if she finds a distribution of fundamentals that rationalize the data, she can be assured that it is the *only* distribution that does so. **Theorem 3** shows that this sort of uniqueness always holds in this model.

**Theorem 3.** *Assume that the researcher observes the values of structural parameters, the matrix of commuting flows  $L(r, s)$ , commuting costs  $\kappa(r, s)$  and trade costs  $\tau(r, s)$ , as well as housing supply  $H(s)$  and wages  $w(s)$  at every location. Then the researcher can recover the set of fundamental amenities  $\bar{a}(r, s)$  and productivities  $\bar{A}(s)$  uniquely (up to scale).<sup>9</sup>*

Although **Theorem 3** guarantees the uniqueness of model fundamentals (conditional on structural parameters and the observed data), it does not offer an algorithm to compute the distribution of these fundamentals, such as the algorithm provided by **Theorem 2**. In fact, one can show that the simple iterative algorithm of **Theorem 2** is not guaranteed to work in general for the model inversion. That said, there exist relatively minor departures from it that can be used. One example is the algorithm used to invert the model in **Desmet et al. (2018)**. This algorithm relies on an approximation to the simple iterative algorithm that is guaranteed to converge to the equilibrium distribution of fundamentals. Moreover, the simple iterative algorithm can be shown to work in certain special cases of the general model. One such example is the special case of the model I consider for the quantitative application of Section 2.2. In that special case, **Section A.4 of the Online Appendix** presents a simple iterative procedure that can be used to invert the model.

2.1.3. *Isomorphisms with other quantitative spatial models*

In this section, I show that the tractability results of Section 2.1.2 carry over to an entire class of quantitative economic geography models that make different assumptions on consumption, production, and market structure. I show this by presenting a set of isomorphisms between the model of Section 2.1.1 and models featuring these alternative assumptions. Finally, I point out which models used in the quantitative economic geography literature belong to this class. The formal proofs of the isomorphisms are relegated to **Section A.3 of the Online Appendix**.

**Land use in production.** I first present an isomorphism between the model of Section 2.1.1 and a model in which the production of tradables requires two inputs: labor and land. The two inputs enter the production function with Cobb–Douglas shares  $\mu$  and  $1 - \mu$ :

$$q(s) = A(s)L(s)^\mu \Lambda(s)^{1-\mu}$$

where  $\Lambda(s)$  is the quantity of land used by the firm. Land is available in fixed positive supply at each location. In equilibrium, land rents adjust at every location until the local land market clears. Rents are redistributed to local employees with equal shares. Thus, the income of an employee at location  $s$  equals her wage income plus a  $1/L(s)$  fraction of rents at  $s$ .

Next, I show isomorphisms between the model of Section 2.1.1 and two models in which locations endogenously specialize in the set of tradable goods they produce. In other words, I relax the assumption

<sup>9</sup> Up to scale means that the researcher cannot recover the worldwide level of either fundamental. For instance, multiplying each location pair’s amenity level by the same constant yields the same equilibrium (except that each worker’s unobserved utility is multiplied by a constant).

that each location produces its own good.

**Endogenous specialization, driven by comparative advantage.** In this model, specialization is driven by differences in comparative advantage, as in **Eaton and Kortum (2002)**. There exists a continuum of tradable goods, indexed by  $\omega \in [0, 1]$ . Workers have constant elasticity of substitution (CES) preferences over tradables:

$$U_i(r, s) = a_i(r, s) \kappa(r, s)^{-1} \left[ \int_0^1 q_i(r, s; \omega)^{\frac{\sigma-1}{\sigma}} d\omega \right]^{\frac{\sigma}{\sigma-1}} H_i(r, s)^{1-\nu}$$

and purchase each tradable good  $\omega$  from the location that offers the good at the lowest price at  $r$  (including trade costs). Immobile local landlords consume tradables only, and have the same preferences over them as workers.

Every location hosts a large number of perfectly competitive firms that can produce any tradable good  $\omega$ . However, locations are not equally well-suited to producing a given good. In particular, the productivity of a good is heterogeneous across locations. As in **Eaton and Kortum (2002)**, I assume that the productivity of good  $\omega$  at location  $s$ ,  $A_\omega(s)$ , is drawn from a Fréchet distribution, independent across goods and locations:

$$Pr [A_\omega(s) \leq z] = e^{-[z/A(s)]^{-\theta}}$$

where  $\theta > \sigma - 1$  is a parameter driving the dispersion of productivity draws, and  $A(s) > 0$  is the productivity of the location. As a result of these assumptions, locations specialize in the set of tradable goods they are most productive at relative to other locations.

**Endogenous specialization, driven by increasing returns.** Finally, I show an isomorphism between the model of Section 2.1.1 and a model in which locations’ endogenous specialization in tradables is driven by increasing returns, as in **Krugman (1991)**. There exists a continuum of tradable goods, indexed by  $\omega \in [0, N]$ , where the set of tradables  $N$  is determined endogenously. As in the previous model, workers have constant elasticity of substitution (CES) preferences over tradables:

$$U_i(r, s) = a_i(r, s) \kappa(r, s)^{-1} \left[ \int_0^N q_i(r, s; \omega)^{\frac{\sigma-1}{\sigma}} d\omega \right]^{\frac{\sigma}{\sigma-1}} H_i(r, s)^{1-\nu}$$

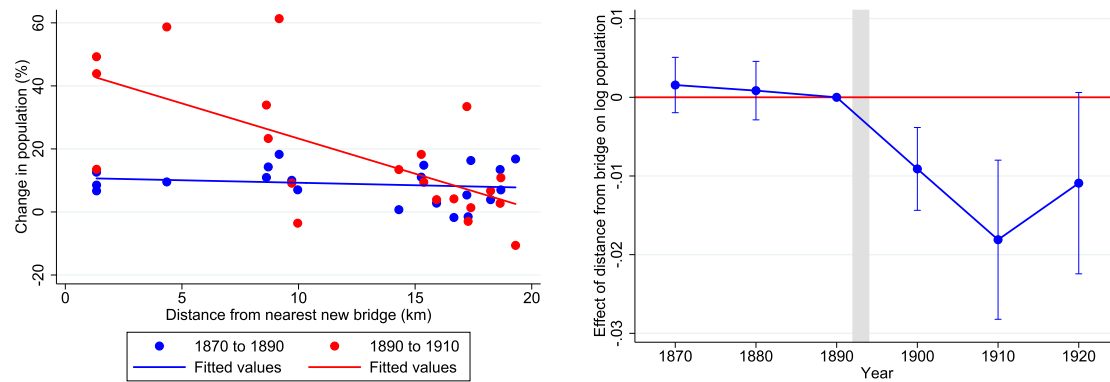
while landlords have the same preferences over tradables.

Firms produce tradable goods with a technology that is subject to increasing returns, as in **Krugman (1991)**. In particular, producing  $q_\omega(s)$  units of good  $\omega$  at location  $s$  requires  $q_\omega(s)/A(s)$  units of labor, plus an additional fixed  $f > 0$  units of labor required for start-up. This latter element of the technology is a fixed cost, which implies that production is subject to increasing returns.

Firms are aware that consumers differentiate across tradable goods. As a result, each good is only produced by one firm in equilibrium, and the firm uses its monopoly power to set a price of its good above the marginal cost. Firms are, however, also aware that they are atomistic relative to the whole set of tradables produced, and therefore cannot influence location-wide prices (monopolistic competition). In equilibrium, each firm operates at its efficient size, implying that its labor demand is a function of local productivity and structural parameters only. Local labor market clearing then determines the equilibrium mass of tradables that a location specializes in.

As a result of these isomorphisms, special cases of the model are isomorphic to a whole range of existing models in quantitative economic geography. In particular,

- the model of Section 2 in **Redding (2016)** is isomorphic to a special case of the model with specialization through comparative advantage in which there are no agglomeration externalities ( $\alpha = 0$ ), no congestion disamenities ( $\lambda = 0$ ) and no commuting;
- the model of Section 3 in **Redding (2016)** is isomorphic to a special case of the model with specialization through increasing returns in which there are no congestion disamenities and no commuting;



**Fig. 1.** The reduced-form effects of bridges on the upper Danube. The left panel is a scatterplot in which each circle represents a settlement in the 20 km radius of bridges that were built on the upper Danube between 1891 and 1895. The horizontal axis corresponds to the settlement's distance from the nearest bridge in kilometers, while the vertical axis corresponds to the settlement's population growth in percentages, either between 1870 and 1890 (blue circles) or between 1890 and 1910 (red circles). The blue and red lines are the corresponding regression lines. The right panel shows the estimated coefficients of column (1) in Table 1 by census year, along with their 95% confidence intervals. The shaded grey area represents the period of bridge construction on the upper Danube. Source: Censuses of the Hungarian Kingdom, 1870, 1880, 1890, 1900, 1910 and 1920. (For interpretation of the references to color in this figure legend, the reader is referred to the Web version of this article.)

- the model of Redding and Sturm (2008) is isomorphic to a special case of the model with specialization through increasing returns in which there are no congestion disamenities, no commuting and no location taste heterogeneity (corresponding to the limiting case of  $\eta \rightarrow 0$ );
- the model of Allen and Arkolakis (2014) is isomorphic to a special case of the model with no housing ( $\nu = 1$ ) and no commuting;
- the model of Monte et al. (2018) is isomorphic to a special case of the model with specialization through increasing returns in which there are no congestion disamenities.

As a consequence of these isomorphisms, the tractability results of Section 2.1.2 carry over to the above frameworks as well, underscoring again that a broad class of quantitative economic geography models exhibit tractability. This is true despite the fact that these models incorporate various dimensions of real-world heterogeneity, such as differences in amenities, productivity, trade costs, and commuting costs across locations. The next section illustrates how this class of models can be used to study important historical questions.

## 2.2. Quantitative illustration: bridges on the Danube

The Danube is Europe's second-longest river, with a total length of 2850 km (1770 mi), roughly identical to the Rio Grande (1759 mi). Over the long segment of the river that was historically located in Hungary,<sup>10</sup> the first permanent bridge was built to connect the western and eastern sides of the country's largest city, Budapest, in 1849. Outside Budapest, however, no bridges were built on the river until the 1890s. Upriver from Budapest, the reason for this was technological: floating ice on the river, a regular occurrence in the spring, would have destroyed any bridge until technology allowed for the construction of more resilient structures in the late 19th century (Tóth, 2009). Once bridge construction on the upper Danube became technologically feasible, three bridges were quickly built on this segment of the river: one in the city of Pozsony (now Bratislava, Slovakia) in 1891, one in the city of Komárom in 1892, and one in the city of Esztergom in 1895.

What was the effect of these new bridges on the spatial distribution of economic activity? How much did they change the costs of trade across locations? Did population relocate in response to the changing trade costs, and if it did, to what extent? How much did the new bridges

increase Hungarian residents' welfare? Some of these questions can be answered using a reduced-form empirical approach that exploits the exogenous, technology-driven timing of bridge construction. I lay out this reduced-form approach in Section 2.2.1. However, for some questions, such as the one about welfare, a quantitative model is necessary. Hence, in Section 2.2.2, I examine the effect of bridge construction in a quantitative framework that belongs to the class of models presented in Section 2.1. Codes that allow for the replication of these exercises are available on the author's website.

### 2.2.1. The reduced-form effects of bridges

In this section, I assess the reduced-form local effects of the new bridges on the upper Danube without committing to a structural model. For this reduced-form assessment, I collect data on the spatial distribution of population within Hungary. Starting from 1870, decadal census data are available on population at the level of *settlements* (település). Settlements are cities, towns or villages. For the censuses of 1870 and 1880, only the population of settlements above 2000 inhabitants is available. Thus, to avoid selection bias, I restrict my reduced-form sample to settlements whose population already exceeded 2000 inhabitants in 1870. I also restrict the sample to settlements in a 20-km radius around the three new bridges. This gives me a sample of 21 settlements, with population observed for all of them in 1870, 1880, 1890, 1900 and 1910, and population observed in 1920 for the nine of them that remained in Hungary after the country's 1920 border changes.<sup>11</sup>

The left panel of Fig. 1 presents a scatterplot of population changes, both before and after bridge construction, against distance from the nearest new bridge for the 21 settlements in the sample. The results are striking. Before bridge construction (i.e., between 1870 and 1890), there was no systematic relationship between distance from the locations of future bridges and settlements' population growth. But once the bridges were present (i.e., by 1910), population grew substantially faster in the bridges' close proximity.

To further investigate the timing of the effects of the new bridges, I run regressions of the form

<sup>11</sup> Restricting the radius to 20 km has the advantage that the 20-km buffers around the three bridges do not overlap. Although census data are also available in the decades after 1920 for the nine settlements that remained in Hungary, the country's extensive border changes would make the interpretation of these estimates difficult. As I argue below, this issue is already apparent for the 1920 estimates.

<sup>10</sup> Nearly one third of the river's entire length was in Hungarian territory before Hungary's large-scale border changes in 1920.

$$\log Pop_{st} = \sum_{y \in \{1870, 1880, 1900, 1910, 1920\}} \beta_y Dist_s \delta_y + \alpha_s + \delta_t + \epsilon_{st}$$

where  $Pop_{st}$  denotes the population of settlement  $s$  in census year  $t$ ;  $\alpha_s$  and  $\delta_t$  stand for settlement and census year fixed effects, respectively; and  $Dist_s \delta_y$  denotes the interaction between distance from the bridge and the fixed effect for census year  $y$ . The coefficient of interest,  $\beta_y$ , measures the effect of distance from the bridge on settlement population in year  $y$  relative to the base year, which I set to 1890.

Column (1) of Table 1 presents the estimated  $\beta_y$  coefficients, which are also plotted in the right panel of Fig. 1. The estimated coefficients are consistent with the pattern suggested by the left panel of Fig. 1. While locations in the new bridges' close proximity did not exhibit different population trends prior to bridge construction, their population grew significantly faster already by 1900, and even more so by 1910. The exogenous timing of bridge construction supports a causal interpretation of these estimates. Column (1) of Table 1 also reports the standardized beta coefficients in curly brackets. These coefficients are large: by 1910, for instance, a one standard deviation decrease in distance from a new bridge led to a 0.123 standard deviation increase in log population.<sup>12</sup>

Columns (2) to (8) of Table 1 present a series of robustness checks to the headline reduced-form finding that the new bridges attracted people to their surroundings, especially by 1910. In column (2), I include a settlement's distance from the country's economic center, Budapest, interacted with census year fixed effects as controls. Next, I include initial settlement population interacted with census year fixed effects to allow for differential population trends across settlements of different size. I conduct this exercise both without (column 3) and with Budapest controls (column 4). Finally, I address the issue that the census only started to include military personnel in settlement populations in 1900. From the two years in which population data are available both with and without military personnel (1900 and 1910), it is clear that the city of Komárom was the only settlement in the sample with substantial military presence. Therefore, columns (5) to (8) repeat the specifications of columns (1) to (4) after dropping Komárom from the sample. The estimated effects of the new bridges on population are largely consistent across these specifications.

The fact that the new bridges drew population to their surroundings should not be surprising. First, there is a large empirical literature documenting that transport infrastructure can attract population and economic activity, either through growth or relocation (Redding and Turner, 2015). Second, Armenter et al. (2014) argue in a spatial model that bridges, in particular, can foster the emergence of cities in their surroundings. The mechanism laid out by Armenter et al. (2014) relies on the fact that a bridge decreases transport costs at a particular point in space along a river. Thus, it makes locations near the river, but especially locations in the bridge's close proximity, see a larger increase in their access to trade, triggering a reallocation of population. This effect may be further amplified by agglomeration economies, which attract even more people to the newly formed city at the bridge. The reduced-form findings of this section can thus be thought of as providing causal evidence for the mechanism laid out by Armenter et al. (2014).<sup>13</sup>

<sup>12</sup> The population effect of the new bridges becomes smaller and turns insignificant at a 5% level by 1920. Hungary's border changes in the same year might be responsible for this fact. In particular, the upper Danube became the border between Hungary and Czechoslovakia, and political conflict between the two countries drastically reduced trade between them (Nagy, 2020b). Hence, the three bridges could no longer facilitate trade in the same way as earlier. Unsurprisingly, estimation precision also drops as population data are only available for nine settlements in 1920.

<sup>13</sup> Another related paper is Tompssett (2020), who finds a positive causal impact of bridges on county populations in the U.S. Midwest.

## 2.2.2. A quantitative model

In Section 2.2.1, I showed that the new bridges on the upper Danube attracted population to their surroundings. But what were the changes in trade costs brought about by the new bridges? As no data on trade costs are available in this historical setting, answering this first-order question about the effects of bridges requires an approach different from reduced-form empirics. In Section 2.1.1, I presented a quantitative spatial model in which trade costs are among the fundamentals shaping the spatial distribution of population and economic activity. It seems reasonable to use this model as a tool to answer how the new bridges reduced trade costs across Hungarian locations. Moreover, a quantitative model also allows one to examine the impact of bridges on aggregate outcomes, such as welfare. Thus, I also use the quantitative framework to study how much the new bridges on the upper Danube increased Hungarian residents' well-being in this section.

Data availability requires me to simplify the structure of the model relative to the general framework of Section 2.1.1. First, commuting data are not available for this period, although it is unlikely that commuting was large-scale at this time in any case. As a result, I consider the version of the model of Section 2.1.1 with no commuting (see footnote 5). Second, while population data are available for the period, wage data are not. Hence, it proves impossible to separate fundamental amenities from fundamental productivity; a location with large population can feature good amenities, high productivity, or both. To solve this issue, I assume that fundamental productivity did not differ across locations and only allow amenities to vary across space. Finally, I abstract from congestion externalities as the model already features dispersion forces in the form of idiosyncratic location tastes and housing. Section A.4 of the Online Appendix outlines the model under these assumptions. While the realism of these assumptions is clearly debatable, one advantage of making them is that they result in an even more tractable model than the one presented in Section 2.1. The existence and uniqueness results of Section 2.1.2 obviously still hold in this special case of the model. At the same time, simple iterative algorithms can be used to both simulate and invert the model in this special case, as Section A.4 of the Online Appendix shows.

I choose the following trade cost function to capture the potential reduction in transport costs brought about by bridges. Between any two locations  $r$  and  $s$  that are on the same bank of the Danube, I assume that the cost of trade was exponential in distance between these two locations,  $dist(r, s)$ :

$$\tau(r, s) = e^{\varphi \cdot dist(r, s)}$$

As a result, parameter  $\varphi > 0$  captures the cost of transporting goods per unit of distance.<sup>14</sup> Between any two locations that are on opposite banks of the Danube, I follow Armenter et al. (2014) and allow for two options. The first option involves crossing the river on a boat at cost  $\psi > 0$ . This amounts to a trade cost between locations  $r$  and  $s$  equal to

$$\tau_{boat}(r, s) = e^{\varphi \cdot dist(r, s) + \psi}.$$

The second option involves crossing the river on a bridge at location  $b$ . Abstracting from the cost of crossing the bridge itself, this amounts to a trade cost equal to

$$\tau_b(r, s) = e^{\varphi [dist(r, b) + dist(b, s)]}.$$

I assume that traders between  $r$  and  $s$  choose the option with the lowest cost, implying that the actual trade cost between  $r$  and  $s$  equals

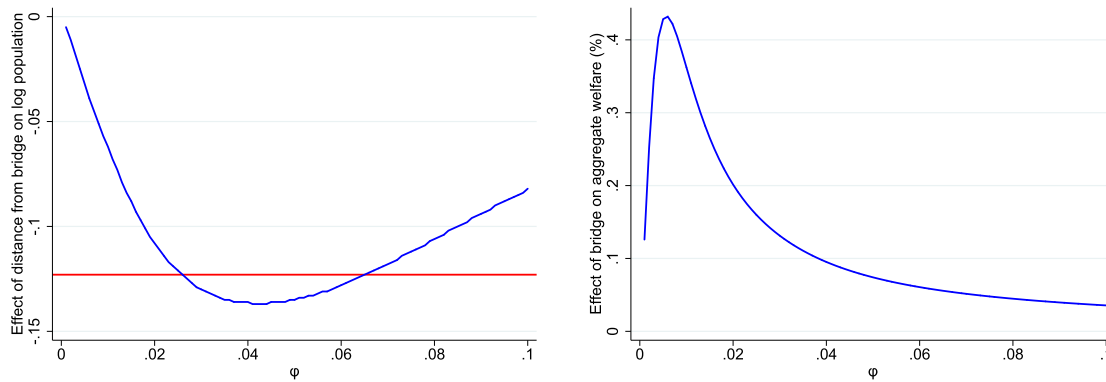
$$\tau(r, s) = \min \left\{ \tau_{boat}(r, s), \min_{b \in B} \tau_b(r, s) \right\}$$

<sup>14</sup> Lacking georeferenced data on transportation networks, I approximate distance between two locations by their distance "as the crow flies." Although a simplification, this assumption is unlikely to lead to large biases since Hungary's railroad network was well-developed by the beginning of the 20th century. In particular, the network was comparable in density to the networks of developed countries (Nagy, 2020b).

**Table 1**  
The reduced-form population effects of bridges.

Dependent variable: Log settlement population								
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Distance from nearest bridge X 1870	0.002 (0.002) {0.011}	0.002 (0.002) {0.011}	0.002 (0.002) {0.013}	0.002 (0.002) {0.013}	0.002 (0.002) {0.016}	0.002 (0.002) {0.017}	0.002 (0.002) {0.018}	0.003 (0.002) {0.018}
Distance from nearest bridge X 1880	0.001 (0.002) {0.006}	0.001 (0.002) {0.006}	-0.001 (0.002) {-0.007}	-0.001 (0.002) {-0.007}	0.002 (0.002) {0.011}	0.001 (0.002) {0.011}	-0.000 (0.002) {-0.002}	-0.000 (0.002) {-0.002}
Distance from nearest bridge X 1900	-0.009*** (0.003) {-0.062}	-0.009*** (0.003) {-0.062}	-0.008** (0.003) {-0.055}	-0.008** (0.004) {-0.056}	-0.008*** (0.003) {-0.055}	-0.007*** (0.002) {-0.054}	-0.007* (0.004) {-0.050}	-0.007* (0.004) {-0.050}
Distance from nearest bridge X 1910	-0.018*** (0.005) {-0.123}	-0.018*** (0.004) {-0.123}	-0.020*** (0.006) {-0.136}	-0.020*** (0.006) {-0.137}	-0.018*** (0.006) {-0.128}	-0.018*** (0.005) {-0.125}	-0.020*** (0.007) {-0.141}	-0.020*** (0.006) {-0.140}
Distance from nearest bridge X 1920	-0.011* (0.006) {-0.052}	-0.011* (0.006) {-0.052}	-0.011** (0.005) {-0.054}	-0.012** (0.005) {-0.055}	-0.010* (0.006) {-0.052}	-0.010 (0.006) {-0.051}	-0.011* (0.005) {-0.055}	-0.011* (0.005) {-0.055}
Settlement FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Census year FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Distance from Bpest X Census yr FE	No	Yes	No	Yes	No	Yes	No	Yes
Log 1870 population X Census yr FE	No	No	Yes	Yes	No	No	Yes	Yes
Drop Komárom	No	No	No	No	Yes	Yes	Yes	Yes
R <sup>2</sup>	0.704	0.723	0.712	0.731	0.676	0.698	0.684	0.706
Number of observations	114	114	114	114	109	109	109	109
Number of settlements	21	21	21	21	20	20	20	20

Each column in this table corresponds to a fixed-effects regression. In all the regressions, the unit of observation is a settlement in the 20 km radius of bridges that were built on the upper Danube between 1891 and 1895, in a given census year (1870, 1880, 1890, 1900, 1910 or 1920). The dependent variable, *Log settlement population*, refers to the settlement's log civil population in years until 1890, and log total (civil and military) population in years from 1900 on. *Distance from nearest bridge* is the settlement's distance from the nearest bridge built on the upper Danube between 1891 and 1895, measured in kilometers. *1870, 1880, 1900, 1910, 1920* refer to census year fixed effects for the specific year. Standardized beta coefficients in curly brackets. Standard errors clustered at the settlement level. \*: significant at 10%; \*\*: significant at 5%; \*\*\*: significant at 1%. Source: Censuses of the Hungarian Kingdom, 1870, 1880, 1890, 1900, 1910 and 1920.



**Fig. 2.** The effects of bridges in the model. The left panel shows the estimated beta coefficients from a fixed-effects regression of log population on distance from the bridge interacted with a 1910 dummy, run on model-simulated data, as a function of transport cost parameter  $\varphi$ . The right panel shows the effect of new bridges on Hungarian workers' expected utility  $\bar{U}$  in the model, as a function of transport cost parameter  $\varphi$ .

where  $B$  denotes the existing set of bridges on the Danube.

I set the boat cost parameter to  $\psi = 126.98\varphi$  based on historical evidence from Fogel (1964). This implies that the cost savings from the bridge depend on the single transport cost parameter  $\varphi$ . Given the central role of this parameter for the research question at hand, I quantify the impact of the new bridges in the model for multiple values of  $\varphi$ . More precisely, I set up a fine grid of  $\varphi$  between 0.001 and 0.1, and repeat the quantitative analysis for each value of  $\varphi$  on this grid.<sup>15</sup>

To measure the impact of new bridges in the model for any given value of transport cost parameter  $\varphi$ , I apply the following strategy. As population data are available at the level of settlements, I define a location as a 1910 Hungarian settlement above 2000 inhabitants.<sup>16</sup> I invert the model to recover Hungarian settlements' fundamental amenities that rationalize these settlements' observed population in 1910. When conducting the inversion, I allow for the possibility of crossing the Danube through any bridge present in 1910 (that is, through bridges in Pozsony, Komárom, Esztergom and Budapest). Next, I simulate the model in the absence of new bridges that were built on the upper Danube between 1890 and 1910 (Pozsony, Komárom and Esztergom), while keeping all other fundamentals fixed. Hence, a comparison of the 1910 equilibrium and the equilibrium without the new bridges reveals the impact of these bridges on the spatial distribution of population and welfare in the model. As discussed, I conduct this comparison for each value of  $\varphi$  between 0.001 and 0.1.

In the left panel of Fig. 2, I present how population reallocates in response to the new bridges in the model. More precisely, I plot the standardized beta coefficient of the regression

$$\log \text{Pop}_{st} = \beta_{1910} \text{Dist}_s \delta_{1910} + \alpha_s + \delta_t + \varepsilon_{st},$$

run on model-simulated data, against transport cost parameter  $\varphi$ . This regression is identical to the one I run to uncover the reduced-form population effects of new bridges in the data (Section 2.2.1). Thus, coefficient  $\beta_{1910}$  captures the causal effect of distance from the bridge on log settlement population in 1910. The estimated standardized beta coefficients vary over a wide range between  $-0.005$  and  $-0.137$ . Unsurprisingly, the relationship between the population effects of new bridges and transport costs is U-shaped. Whenever transport costs are low or prohibitively high, new bridges have little impact on most locations' access to trade, and hence lead to little reallocation in population.

<sup>15</sup> I set the values of the model's remaining structural parameters to central values used in the literature. In Section A.5 of the Online Appendix, I discuss these parameter choices in detail.

<sup>16</sup> This amounts to inverting and simulating the model for 1801 locations. For simplicity, I abstract from international trade between Hungarian settlements and the rest of the world.

In line with this logic, I also find that the model-implied effect of new bridges on Hungarian residents' welfare is hump-shaped in transport costs (right panel of Fig. 2). Under low or high transport costs, bridges have little impact on trade, and hence on aggregate welfare. Under intermediate values of transport costs, the aggregate welfare effect of new bridges reaches 0.2%–0.4%, which is far from negligible.

It is straightforward that different values of  $\varphi$  have different implications on the cost reductions brought about by bridges as well. How to pick the “right” value of  $\varphi$ ? One potential strategy is matching the estimated causal effect of bridges between the model and the data. Recall that the standardized beta coefficient of the population effect by 1910 equaled  $-0.123$  in the data (Section 2.2.1). In the left panel of Fig. 2, the horizontal line corresponds to this value. As can be seen, there are two values of  $\varphi$  at which this reduced-form coefficient is matched by the model (0.026 and 0.065). What are the cost reductions implied by these values of  $\varphi$ ? Obviously, such cost reductions vary across space, but one can calculate them for any specific settlement pair. For instance, if  $\varphi = 0.026$ , then trade costs between Győr and Érsekújvár, two relatively large cities on opposite banks of the upper Danube, reduced by 95% once they had the opportunity to trade through the new Komárom bridge. The corresponding cost reduction estimate is 99.4% if  $\varphi = 0.065$ .

These cost reductions, even the more modest of the two, seem unrealistically large. In fact, Nagy (2020b) estimates per-kilometer transport costs from spatial price differences of traded commodities in 1910 Hungary, and finds a much smaller value of  $\varphi = 0.000351$ . These findings underscore that quantitative modeling always needs to be accompanied with caution and care. Instead of concluding that trade costs in early-20th century Hungary were orders of magnitude higher than previously thought, the researcher should ask a series of questions when confronted with these findings. For realistic values of transport costs, why does the model imply a smaller population effect of bridges than the data? What forces are missing from the model that might be amplifying the population-attracting effect of bridges in reality? In the next section, I discuss a few forces that are missing from the class of models presented in Section 2.1 and might be responsible for this amplification. At the same time, I discuss how incorporating these forces with the aim of further realism can prove challenging from a tractability perspective.

### 2.3. Additional model ingredients: challenges to tractability

In this section, I consider a set of model ingredients that are absent from the class of quantitative models discussed in Section 2.1. My primary focus is on how these additional ingredients can pose challenges to tractability, and how existing quantitative studies have addressed these challenges.



**Dynamics.** Many historical questions are dynamic by nature: How did transport infrastructure foster economic growth? How persistent has the spatial distribution of economic activity been? Answering these questions with a quantitative spatial model necessarily requires incorporating dynamics in the model. There are also other cases in which, even though the question the researcher is after is not inherently dynamic, dynamics play a quantitatively relevant role. In the application of Section 2.2, recall that the reduced-form effect of new bridges seemed to increase over time, at least before 1920 (right panel of Fig. 1). This could suggest that a bridge not only had a static impact on the economy of its surroundings, but also fostered local growth. As such growth effects are not present in the model, this could explain why the model underestimates the effects of bridges under realistic levels of transport costs.

It is not hard to see why incorporating dynamics can be challenging from the point of view of model tractability. In a model in which the future enters economic agents' objective functions and influences the decisions they make today, agents need to predict the future evolution of prices. In a model with spatial linkages across locations, the problem is even more complex: agents need to predict the evolution of prices *at each location*. Without simplifying the problem, the curse of dimensionality makes it infeasible to compute the equilibrium, even if the number of locations is small.

Various simplifications have been offered to this general dynamic problem. One possible simplification is, of course, assuming that agents do not care about the future and thus make static decisions at every point in time. Under this assumption, the model still needs to be solved for every time period, but it can be solved as a sequence of static problems. A case in which such an assumption is justifiable is one in which a time period is about as long as a person's lifetime (Delventhal, 2018). A similar assumption can be made if agents live for two time periods, but their decision with dynamic consequences is made in only one of those two periods (Allen and Donaldson, 2020). In the case of firms, a possible microfoundation for the assumption that agents do not care about the future is that irrespectively of their decisions today, their future profits are driven down to zero by free entry (Desmet and Rossi-Hansberg, 2014; Desmet et al., 2018; Nagy, 2020a).<sup>17</sup> Another simplifying assumption can be the absence of trade costs, in which case prices equalize across locations and agents only need to predict the future evolution of one worldwide price (Eckert and Peters, 2018).

**Multiple sectors.** The class of models discussed in Section 2.1 does not feature multiple sectors. From the point of view of the application of Section 2.2, this might be an important omission. In particular, Nagy (2020a) argues in the context of 19th-century U.S. railroad construction that incorporating a rural ("farm") and an urban ("non-farm") sector in a quantitative spatial model dramatically alters the evaluation of transport infrastructure improvements. The intuition for this result is as follows. In Nagy (2020a), urban activities are more subject to increasing returns than rural activities. Therefore, if transport infrastructure improvements foster the growth of urban locations, they can boost these increasing returns and have an impact on the economy that exceeds the savings from transport cost reductions. Similar forces may have operated in the historical context of Section 2.2, another potential reason why the model stops short at explaining the reduced-form effect of bridges in its entirety.

Various other historical questions call for incorporating multiple sectors in the model as well. This is particularly true for questions

related to structural transformation (Delventhal, 2018; Eckert and Peters, 2018; Fajgelbaum and Redding, 2018). Although the tractability properties of Section 2.1 carry over to multi-sector models under certain restrictive assumptions (constant expenditure shares, no input-output linkages across sectors, no agglomeration externalities – Allen and Arkolakis, 2014), such assumptions prove too restrictive in the context of structural transformation. One of the major empirical facts documented about structural transformation, for instance, is that sectoral expenditure shares do change over time (Herrendorf et al., 2014). As a result, these studies cannot rely on the theoretical tractability results of Section 2.1.

**Endogenous infrastructure development.** Another assumption of the class of models discussed in Section 2.1 is that trade and commuting costs are exogenous. Thus, the development of transport infrastructure can be fed into these models as an exogenous reduction in trade (and possibly commuting) costs across certain locations. This methodology cannot account for the fact that infrastructure development is endogenous: whether to develop infrastructure across certain locations is a decision made by agents who take into account the benefits and costs of infrastructure development. In the context of the quantitative application of Section 2.2, it could be the case that new bridges fostered the development of other infrastructure, such as railroad or road connections to settlements where the new bridges were located. If this was the case, then settlements near the new bridges would have gained not only from the bridge, but also from this complementary infrastructure. The fact that the model does not account for such a mechanism can be yet another source of the discrepancy between the effect of bridges in the model and their reduced-form effect in the data.

Studying the forces behind endogenous infrastructure development decisions is the focus of various quantitative historical studies (Santamaria, 2020; Swisher, 2014; Trew, 2020). However, without further assumptions, the problem of endogenous infrastructure development is subject to a similar curse of dimensionality as the dynamic problem discussed earlier in this section. The reason for this is simple: the number of possible links across locations increases exponentially with the number of locations. Therefore, the problem quickly gets computationally out of hand as the number of locations becomes large. A further complication is that infrastructure developers might engage in strategic interaction with one another (Swisher, 2014).

Simplifications to the problem offered in this literature rely on reducing the number of links and restricting strategic behavior (Swisher, 2014), assuming one-dimensional space and free entry in infrastructure development (Trew, 2020), or making infrastructure development location-specific rather than link-specific (Santamaria, 2020; Fajgelbaum and Schaal, 2020; Ducruet et al., 2020). That said, even these simplifying assumptions do not necessarily guarantee certain aspects of tractability, such as equilibrium uniqueness. Facing this issue, Trew (2020) assumes that the infrastructure allocation maximizing economy-wide net rents is selected by the U.K. Parliament in the case of multiple equilibria with different allocations.

As discussed above, the tractability issues arising in quantitative models with such additional ingredients often prove challenging and require the researcher to make highly restrictive assumptions. This, of course, raises the question of whether quantitative modeling is the way to go to answer these important historical questions. However, one needs to weigh these costs of quantitative models against their benefits. There are at least four benefits that need to be taken into account. First, quantitative models can be used to infer missing data, as already highlighted in the application of Section 2.2 and discussed further in Section 3. Second, unlike reduced-form empirical techniques, they are able to measure the aggregate general equilibrium impact of historical events and therefore distinguish growth from reallocation. Third, those quantitative models that feature dynamics can be used to study the long-run effects of events. Due to various other events taking place over long time periods, such long-run effects are often impossible to recover empirically.

<sup>17</sup> These papers allow firms to innovate, which is a decision with dynamic consequences as it increases the firm's future productivity. However, they assume that innovation spills over locally after one period, and new firms can freely enter at any point in time. Hence, entry drives down firms' profits to zero in all future periods, and they only innovate to the extent that maximizes their current profits – a static decision. In Desmet et al. (2018), a similar set of assumptions is made on individuals' migration costs, which guarantees that their location choice also reduces to a static decision.

cally.<sup>18</sup> Finally, issues of tractability are ubiquitous in structural modeling and are by no means restricted to the field of economic geography. Therefore, new techniques developed to address tractability issues may spill over and benefit other areas of economics as well.

### 3. Availability of historical data

Available data become sparser as one goes back in time. This is by no means a problem that only economists face. In archaeology, conclusions normally need to be drawn from scattered, small remains. Identifying the sex and age of skeletal remains, for instance, relies primarily on examining the pelvis, the skull, and other large bones. Quite often, such remains are not available. In that case, the calcification of teeth and the maturation of non-cranial bones may provide information on age, while comparing the two may be informative about sex (Ubelaker, 2008). However, even these might be absent. What is left of the entire pre-human species *Australopithecus bahrelghazali*, for example, is three partial jawbones and a tooth (Brunet, 2010).

Similarly, economic geographers studying historical questions need to draw conclusions from a small amount of sparse data, which become increasingly sparser as one goes back to the more and more distant past. This problem is, if anything, more severe in the case of quantitative economic geography studies. This is because taking quantitative models to the data typically requires observing a special set of data – namely, macroeconomic aggregates such as GDP, population or land values at the location level. In what follows, I briefly review the availability of historical data sources on GDP, sectors and occupations, land values, population, transportation networks, commuting and geographic characteristics that have been used in the quantitative economic geography literature. Next, I discuss what Geographic Information Systems (GIS) and structural modeling can add to the information embedded in these sparse data.<sup>19</sup>

**GDP.** National accounting is almost exclusively a 20th-century phenomenon. Official annual estimates of national income were first compiled by the Australian government in 1886, followed by Canada in 1925 (Bos, 1992). With Simon Kuznets joining the National Bureau of Economic Research in 1929, the United States soon became the leader in preparing national income estimates. In 1939, the U.S. Commerce Department produced the first estimates of income at the state level (Carson, 1975). However, it was not until 1947 that the first international guidelines on national accounting were prepared (Bos, 1992). Thus, official data on GDP are virtually unavailable for the study of historical questions that predate the mid-20th century.

That said, economic historians have produced estimates of past GDP based on available population, income and production data. One notable example is the Maddison database, which includes estimates of country GDP going back to the year 1 (Bolt et al., 2018), although it has received criticism for its lack of accuracy (Fariss et al., 2017). In the case of the U.S., Easterlin (1960), Gallman (1966), David (1967) and Weiss (1992) estimate per capita income of the country and its regions for various overlapping periods during the 19th century, relying primarily on sectoral data. The vast differences across their estimates already suggest that they are likely subject to large measurement errors. Hence, using them as direct inputs into quantitative models could cast doubt on the model's quantitative predictions. Nevertheless, such estimates of historical GDP might still be suitable for testing the model's qualitative

<sup>18</sup> By identifying these long-run effects, historical studies also have the potential to highlight economic mechanisms that require a very long time to unfold. An example is the century-long change in occupational structure due to improvements in communication technologies (Michaels et al., 2019).

<sup>19</sup> An issue related to data scarcity is that historical data, even if available, are often stored in archives in a not (yet) digitized format. Combes, Gobillon and Zylberberg (2020) discuss recent machine learning techniques aimed at processing such data, along with the associated challenges.

predictions. For example, Nagy (2020a) finds that, according to a spatial model of the 19th-century U.S. economy, the U.S. Northeast had higher per capita income than the South and the Midwest before the Civil War. This pattern of income distribution across U.S. regions is in line with the above historical estimates. Such tests, even though they do not force us to fully trust historical estimates of GDP, can lend additional credibility to models that can replicate similar broad patterns observable in the estimates.

**Sector- and occupation-level data.** The 19th century saw the first sectoral censuses conducted. These censuses tend to provide data on employment, the number of establishments, land use, and/or output in certain important sectors such as agriculture or manufacturing. Quite often, however, they do not cover the entire economy.

In the United States, James Madison suggested that the first, 1790 population census include occupational statistics. Though this did not happen in the end, the 1810 census asked enumerators to collect data on the country's manufacturing establishments. This became the first U.S. Census of Manufactures, and it was followed by similar censuses conducted in 1820 and 1840. However, the quantity and quality of the collected data in these first three censuses disappointed contemporaries.<sup>20</sup> Major improvements were made afterwards, which made the 1850 and subsequent decennial Censuses of Manufactures substantially more reliable than the previous ones (Fishbein, 1973). In the meantime, the first U.S. Census of Agriculture was conducted in 1840,<sup>21</sup> and was followed by decennial censuses of this sector since then. The quantitative economic geography literature has used these historical Censuses of Agriculture and Manufactures. In particular, Donaldson and Hornbeck (2016) use Census of Agriculture data to measure the effect of late-19th century railroads on the value of agricultural land. Eckert and Peters (2018), on the other hand, use Census of Manufactures data from 1880 on to study the spatial patterns of structural change towards manufacturing.

Another country with a long history of sectoral employment data is the United Kingdom. The 1801 U.K. census already included a question about occupation, albeit only at a broad level of sectoral disaggregation. Detailed census data on occupation are available from 1851 on. Moreover, Shaw-Taylor et al. (2010) compile data on the occupation of males in 1710 (for England) and 1817 (for England and Wales) based on thousands of baptism records. The occupational categories are such that the sector can be inferred from them; examples of occupations are “grower of minor crops,” “coal miner,” “textile products maker” and “rag dealer.” Trew (2014) and Trew (2020) use these occupational data at the level of *registration districts*<sup>22</sup> to study spatial structural change in England during the industrial revolution.<sup>23</sup>

**Value of land and structures.** The value of land and buildings on farms were included in the U.S. Census of Agriculture from 1850 on. Naturally, these reflect not only land values but also the value of capital embedded in buildings. To separate the two, Fogel (1964) estimates the value of agricultural land alone at the state level. Donaldson and Hornbeck (2016) use these estimates to impute the value of agricultural land at the county level, assuming that the value of buildings relative to land was uniform within states.

Historical data on the value of non-agricultural land and structures are also sparse. In England and Wales, however, the value of proper-

<sup>20</sup> One reason for this is that the instructions given to enumerators were vague, especially in 1810. Another reason is that several establishments refused to provide data, fearing that the information would be used for tax purposes.

<sup>21</sup> The 1840 census already included detailed data on livestock, prices and output by crop, as well as estimates of the value of agricultural production by county (U.S. Census, 1841).

<sup>22</sup> The average registration district in England and Wales had 28,700 inhabitants in 1851.

<sup>23</sup> Further quantitative economic geography papers use historical sector-level data from Argentina (Fajgelbaum and Redding, 2018), Germany (Peters, 2019) and India (Donaldson, 2018).

ties were constantly assessed by the government since 1601 for tax purposes. These so-called *rateable values* again reflect both the value of land and the value of structures built on it. To study the economic impact of the steam railway on London, [Heblich et al. \(2020a\)](#) use these rateable values by borough in the Greater London area. In the case of Berlin, [Ahlfeldt et al. \(2015\)](#) use high-resolution data on the value of undeveloped land (i.e., without structures) collected by a government-appointed building surveyor ([Kalweit, 1937](#)) to examine the effect of the Berlin Wall on the city's economy.

**Population.** Due to the long history of government-administered population censuses, population data tend to be the most widely available and the highest in quality.<sup>24</sup> As a result, these data have been widely used in the quantitative economic geography literature. Historical U.S. population data at the county level are used in [Allen and Donaldson \(2020\)](#), [Desmet and Rappaport \(2015\)](#), [Donaldson and Hornbeck \(2016\)](#), [Eckert and Peters \(2018\)](#) and [Nagy \(2020a\)](#), among others. Census-based population data at similar levels of geographic disaggregation are used in quantitative historical studies on Argentina ([Fajgelbaum and Redding, 2018](#)), Germany ([Peters, 2019](#); [Santamaria, 2020](#)) and the United Kingdom ([Trew, 2014, 2020](#)). Normally, census data also allow for distinguishing urban from rural populations. This has been an advantage for papers investigating the historical drivers of urbanization ([Fajgelbaum and Redding, 2018](#); [Nagy, 2020a, 2020b](#); [Redding and Sturm, 2008](#)).

Historical census data can also be used to examine the drivers of the distribution of population within cities. Within-city analysis necessitates a distinction between *residential* population (i.e., number of people by location of residence) and *workplace* population (i.e., number of employed people by workplace location). Data on the former are more widely available, as censuses are normally conducted on residential locations. [Heblich et al. \(2020a\)](#), for instance, use residential population by borough in Greater London from 1801 until 1921 to document that the steam railway allowed people to move away from downtown areas. To study the effect of the Berlin Wall, [Ahlfeldt et al. \(2015\)](#) use street-level residential population data on Berlin from the 1933 census. They construct a proxy of workplace population by census block by combining district-level employment data with a registry of company locations within districts ([Ahlfeldt et al., 2015](#)).

Population censuses often include data on socioeconomic status such as education, occupation and – more recently – income. As a result, they can also be used to examine the drivers of residential segregation. Various studies have documented that residential segregation patterns exhibit persistence over time. That is, neighborhoods that were poor in the past also tend to be poor today. Multiple explanations have been offered for this phenomenon, based on persistent natural amenities ([Lee and Lin, 2018](#)), differences in 19th-century pollution across Western and Eastern parts of cities ([Heblich et al., 2020b](#)), and the war-time bombing of certain neighborhoods ([Redding and Sturm, 2016](#)).

Finally, certain research questions ask for population data that predate the official censuses. Similar to GDP, the Maddison project estimates country populations back to the year 1 ([Bolt et al., 2018](#)). At a higher level of spatial disaggregation, the *History Database of the Global Environment* (HYDE) combines an assortment of historical sources to estimate the population of each 5 by 5 arc minute grid cell of the Earth, going back 12,000 years ([Klein Goldewijk et al., 2010](#)). [Delventhal \(2018\)](#) uses these data for the year 1000, aggregated up to a 3° by 3° resolution, to calibrate a dynamic spatial model of the world economy. He uses the model to quantify the extent to which falling trade costs in the last 1000 years contributed to the current world income distribution.

<sup>24</sup> This does not mean that population data cannot be patchy sometimes. 1850 census data on San Francisco were destroyed in a fire, while data on Contra Costa and Santa Clara counties were lost on the way to the San Francisco office ([U.S. Census, 1852](#)).

**Transportation links.** A substantial portion of the quantitative economic geography literature has pointed out the importance of spatial linkages across locations. First, spatial linkages might be crucial in transmitting the effects of historical events from locations directly affected by the event to other locations in the economy. Second, the historical event that the researcher aims to evaluate might be the establishment of a spatial linkage itself. For instance, a large literature has studied the general equilibrium effects of U.S. railroads, starting with the seminal work of [Fogel \(1964\)](#).<sup>25</sup> Conducting such an exercise requires collecting a database of transportation links across locations. Such a database essentially consists of a series of maps, one by time period and mode of transportation (such as water, rail and road; or, in the case of within-city transportation, walking, car and public transport).<sup>26</sup> Using data on the cost of transportation by mode, the next task is determining the overall transport cost between any pair of locations in each time period. This is normally done using least-cost algorithms (Dijkstra algorithm, Fast Marching algorithm).

Some studies use data on the location of planned, rather than actual, transportation links. For example, [Santamaria \(2020\)](#) uses planned highways in 1930s Germany to calibrate an economic geography model with endogenous infrastructure placement. [Brinkman and Lin \(2019\)](#) use planned freeways as instruments for the location of actual freeways, which are endogenously affected by opposition to construction due to freeway disamenities.

**Commuting and other travel flows.** Commuting data are useful to calibrate quantitative models of within-city geography, as such data create a mapping between residential and workplace locations. [Heblich et al. \(2020a\)](#), for instance, observe 1921 commuting data across boroughs in the city of London. They combine these data with population by borough to recover the location of people's workplaces in London. This is an important contribution as workplace data within cities are not available in most countries prior to the second half of the 20th century. Unfortunately, such comprehensive commuting data are rarely available either when it comes to other historical contexts. However, it is sometimes possible to rely on survey data. For example, [Brinkman and Lin \(2019\)](#) use data from 1950s travel surveys for Chicago and Detroit to show that freeways acted as barriers to travel flows aimed at the opposite side of the freeway.<sup>27</sup>

**Geographic characteristics.** It should come as no surprise that various studies in quantitative economic geography rely on data on natural geographic characteristics. First, natural features such as rivers and oceans had an influence on transportation routes, particularly during historical times. Thus, they serve as important inputs to the creation of transportation links databases ([Delventhal, 2018](#); [Donaldson and Hornbeck, 2016](#)). Second, nature matters for locations'

<sup>25</sup> Fogel's pioneering idea was that the aggregate impact of railroads on U.S. agriculture needs to be assessed by comparing the real-world economy to a hypothetical economy in which the transportation of agricultural goods can only happen through other modes – an exercise that today's economists call a *counterfactual*. Motivated by Fogel's analysis, [Donaldson and Hornbeck \(2016\)](#) conduct the same exercise in a quantitative economic geography model. Similar to Fogel, they find that the effect of railroads on U.S. agriculture was rather modest. However, [Pérez-Cervantes \(2014\)](#) finds, conducting a large number of counterfactuals involving alternative railroad networks in a similar model, that certain rail connections were much more influential than others. Finally, using a dynamic spatial model of city formation, [Nagy \(2020a\)](#) argues that, even with a limited effect on the agricultural sector itself, railroads had a large aggregate impact on the U.S. economy by fostering the development of cities.

<sup>26</sup> Historical maps are often georeferenced to modern maps when creating such transportation links databases. Unfortunately however, historical maps often lack precision, an issue discussed in more detail in [Hanlon and Heblich \(2020\)](#).

<sup>27</sup> Historical trade data at a higher spatial resolution than cross-country flows are even harder to find. To study the role of trade in Bronze Age city formation, [Barjamovic et al. \(2019\)](#) use the number of mentions on clay tablets as a proxy of trade flows across ancient Assyrian cities.

agricultural potential. High-resolution spatial datasets on agricultural potential include the Food and Agriculture Organization's Global Agro-Ecological Zones database (FAO GAEZ), the Caloric Suitability Index (Galor and Özak, 2016), and the index of agricultural land suitability built by Ramankutty et al. (2002). These datasets are used as inputs to quantitative historical studies such as Delventhal (2018) and Nagy (2020a).

Spatial data are often processed in Geographic Information Systems (GIS). A Geographic Information System is a versatile computer-based tool that helps with the "storage, analysis, output, and distribution of spatial data and information" (Bolstad, 2005). As the quote highlights, GIS is useful for at least three purposes. First, it allows for the efficient storage of high-resolution spatial data. Second, it can perform various sorts of analysis on spatial data, such as merging data up to higher units of disaggregation, measuring distances across spatial units, creating distance buffers around them, and so on. Finally, GIS is especially well-suited to creating nice visual output of spatial data, such as maps. Due to these advantages of GIS, various historical spatial datasets are primarily made available in GIS-compatible format. For instance, the *National Historical Geographic Information System* (NHGIS) is a free online source of GIS-compatible U.S. census and survey data that go back to 1790 (Manson et al., 2017).

Needless to say, no processing tool can compensate for the sparse nature of historical spatial data. That said, data requirements are sometimes less demanding in historical contexts. For instance, commuting was not a widespread phenomenon before the mid-19th century (Heblich et al., 2020a). As a result, studies looking at earlier periods are freed from the need to acquire commuting data. Similarly, population can be a good proxy of income in Malthusian environments.<sup>28</sup> And even if these conditions are not met, another advantage of quantitative modeling is that the model can substitute for missing data.

One notable example of the model substituting for missing data is Barjamovic et al. (2019). They combine the extremely patchy data available from ancient Assyria (a proxy of trade flows across Assyrian cities and the locations of cities that have been found by archaeologists) with a quantitative trade model to predict the locations of cities that have not yet been found. A general issue with this approach is that the structure of the model, which identifies the unobserved data, is an assumption that is untestable in practice. Nonetheless, even though one cannot test the structure of the model in its entirety, one can test some of its implications on observed outcomes that are not used in the model's quantification. For example, Barjamovic et al. (2019) contrast their model's predictions on merchants' itineraries to actual itineraries found. Although they do not use the actual itineraries in the quantification, they find a good fit between the model-implied and the actual itineraries. Moreover, they find that their location predictions for lost cities often coincide with archaeologists' conjectures, and that their model-based method of locating cities works well when predicting the locations of known cities. Such overidentification tests are crucial to show that the model is a credible enough tool to fill the gaps in sparse historical data.

#### 4. Identification

One of the largest challenges facing any empirical investigation in economic geography is the identification of causal effects (Redding and Turner, 2015). In the case of quantitative economic geography studies, this challenge can arise in two separate places. First, the challenge of identification is present when the quantitative model is estimated or calibrated to data. Second, besides taking the model to the data, a large number of quantitative studies also estimate the reduced-form effects of spatial events on local outcomes, as also illustrated in the quantitative

example of Section 2.2. In this section, I focus on the first issue, as it is the one directly related to quantitative modeling. For a review of identification challenges faced by reduced-form empirical work in economic geography, see Baum-Snow and Ferreira (2015).

The key challenge involved in identifying the effects of historical spatial events stems from the fact that the event itself is often endogenous. Transport infrastructure construction, trade, or the spatial concentration of population are the outcomes of decisions made by economic agents. A model in which they occur exogenously misses a set of structural equations that determine these outcomes as a function of other economic variables. Taking the model to the data without these missing equations can bias the estimation or calibration of model parameters entering the equations that are present in the model. A related issue is that of omitted variables. If the set of locations directly affected by the historical event are special in other respects, then the estimation or calibration might falsely attribute the effects of these omitted variables to the event itself.

Quantitative historical studies have developed three broad strategies to overcome the challenge of identification. Some studies rely on natural experiments to estimate the parameters of the model. Others use instrumental variables that are arguably exogenous. Finally, some studies identify the model's parameters using data that only come from the period before the event occurred. In what follows, I briefly review the studies that have followed these three strategies.

**Estimation of model parameters using natural experiments.** A number of quantitative studies exploit natural experiments. In a spatial context, a certain event constitutes a natural experiment if locations' exposure to it is as good as random. That is, which locations are exposed to the event and to what extent are not related to these (or other) locations' pre-event characteristics. The redrawing of borders and expulsion of populations that follow was often provide such natural experiments. Ahlfeldt et al. (2015), for example, use the division of Berlin after the Second World War as a natural experiment to study the strength of agglomeration economies within the city. The division of Berlin ultimately resulted in the construction of the Berlin Wall, which cut the Western part of the city from the Eastern part. Western locations close to the wall experienced a dramatic decline in the surrounding concentration of population and economic activity. The decline was naturally smaller for locations farther away from the wall. As the location of the wall was chosen based on non-economic (military) considerations, the extent to which different locations suffered a decline in surrounding concentration was *as if* these declines were randomly assigned to locations. Thus, Ahlfeldt et al. (2015) can use the division of Berlin as a natural experiment to estimate the model's parameters driving the extent to which locations benefit from surrounding concentration (agglomeration economies).

Ahlfeldt et al. (2015) estimate their quantitative spatial model of Berlin using Generalized Method of Moments (GMM). This method relies on setting up a series of *moment conditions*, i.e., conditions that need to be satisfied by the estimated parameters of the model. Ahlfeldt et al. (2015) base these conditions on the natural experiment provided by the division of the city. In particular, these moment conditions state that a location's proximity to the Berlin Wall was unrelated to the change in the levels of fundamental amenities and productivity at the location. This condition must be satisfied if the placement of the wall was truly exogenous.

Nagy (2020b) and Peters (2019) employ similar strategies to look at the effect of trade on urbanization and the effect of population on local economic growth, respectively. In Nagy (2020b), the natural experiment stems from the redrawing of Hungary's borders after the First World War. This redrawing of borders disproportionately reduced the trading opportunities of regions close to the new border. Nagy (2020b) develops a quantitative model in which trading opportunities affect urbanization, and estimates the key parameter of the model using the variation in changing trading opportunities due to different regions

<sup>28</sup> See Delventhal (2018) for a dynamic spatial model in which the world endogenously transitions out of a Malthusian steady state.

being differentially exposed to the new border. In the case of Peters (2019), the natural experiment comes from the expulsion of ethnic Germans from Eastern European countries after the Second World War. These refugees were settled in Germany, increasing local population substantially but differentially across locations. As the placement of the refugees was driven by non-economic considerations, the variation in population increases across locations can be used to estimate the parameters of a quantitative model in which local population drives local growth.

#### Estimation of model parameters using instrumental variables.

Instrumental variables provide an alternative to the use of natural experiments, which are rare in history. In the context of measuring the effects of spatial events, an instrumental variable is a variable that is correlated with locations' (possibly endogenous) exposure to the event, but does not affect locations' economic outcomes in other ways. Heblich et al. (2020b), for instance, study the persistence of neighborhood segregation in U.K. cities using a dynamic model of residential sorting. A key structural equation of the model is one that links current to past segregation. Estimating the equation with ordinary least squares would be subject to the issue that past segregation is endogenous. As a result, Heblich et al. (2020b) use 19th-century pollution as an instrument for past segregation. This instrument has a significant effect on past segregation, as richer residents sorted to less polluted neighborhoods historically. This method of identification is valid as long as variation in 19th-century pollution across neighborhoods only affects current segregation through segregation in the past. Although this assumption is untestable by definition, Heblich et al. (2020b) also show that a key determinant of 19th-century pollution was whether a neighborhood is on the Eastern or Western side of a city (as winds typically blow from West to East). Thus, a substantial part of the identifying variation comes from geography, which is exogenous.

Various other quantitative historical studies rely on geography-based instruments to estimate the parameters of quantitative models. Heblich et al. (2020a) use straight-line distances across boroughs to instrument rail travel times in early-20th century London. Fajgelbaum and Redding (2018) use least-cost paths to major ports and Spanish colonial postal routes as instruments for the location of railroads in Argentina. In Brinkman and Lin (2019), planned freeways and historical exploration routes serve as instruments for the location of actual freeways in the United States.

**Calibration of the model to pre-event data.** In a few studies, the quantitative model is calibrated to data that precede the event whose effect the model is used to measure. As a result, the potential endogeneity of the event cannot bias the identification of the model's parameters. For example, Nagy (2020a) evaluates the effect of early-19th century U.S. railroads on city development and aggregate growth. Though the placement of railroads is likely endogenous, the calibration of the model is fully based on data from the period prior to railroad construction. One concern could still be that people anticipated the placement of railroads at certain locations even before they started to be built. However, this is unlikely in the case of the steam railroad, which came into existence and spread out from Britain over the course of a few years in the 1820s and 1830s. Other quantitative historical studies calibrating model parameters to pre-event data include Redding and Sturm (2008) and Santamaria (2020), who take their models to pre-WWII German data to study the impact of the 1947 division of Germany on city populations and highway construction, respectively.

As in the case of new techniques meant to increase model tractability, novel ways of identification developed in the quantitative historical geography literature may benefit other researchers. Historical instruments, for instance, are often useful to obtain exogenous variation in present-day outcomes. Among others, Duranton et al. (2014) and Duranton and Turner (2011) use historical exploration routes as instruments to study how the current U.S. road network affects trade and traffic congestion, respectively.

## 5. Conclusion

Over the last few years, the use of quantitative models has gained prominence in economic geography. In this article, I have reviewed the part of this literature that studies the economic impact of historical events with a quantitative methodology. I have argued that this methodology is able to bridge the gap between reduced-form empirical work and classical structural modeling. On the one hand, it is suitable to be combined with rich spatial data, thus allowing the data to speak about important historical questions. On the other hand, structural by nature, quantitative modeling can measure the aggregate impact of historical events and distinguish growth from the mere relocation of economic activity. These benefits from quantitative modeling, however, come at a price. Specifically, studying historical spatial questions with quantitative models is associated with three key challenges: model tractability, the sparse nature of historical data, and identification issues. In this article, I have reviewed how the literature has addressed these challenges.

Even though quantitative historical studies have come a long way over the course of a few years in economic geography, there is clearly room for further research in the field. As already mentioned, historical questions are often important for today's economists because similar events are expected to take place in the future. Developing models that can be used by policymakers to assess the economic impact of these future events has the potential to inform these policymakers' decisions. Quantitative models of economic geography are still rarely used in policy work. One reason for this might be the complex structure of these models, which often makes them seem like a black box to researchers outside the literature. Simplifying and clarifying model structure, without giving up too much on model realism, seems a fruitful direction that may allow these models to be adapted to a larger extent in the world of policy.

#### Declaration of competing interest

The author declares that he has no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

#### Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.regsciurbeco.2021.103675>.

#### References

- Acemoglu, D., 2009. *Introduction to Modern Economic Growth*. Princeton University Press.
- Ahlfeldt, G., Redding, S., Sturm, D., Wolf, N., 2015. The economics of density: evidence from the Berlin wall. *Econometrica* 83 (6), 2127–2189.
- Allen, T., Arkolakis, C., Li, X., 2020. On the Equilibrium Properties of Network Models with Heterogeneous Agents. Mimeo.
- Allen, T., Arkolakis, C., 2014. Trade and the topography of the spatial economy. *Q. J. Econ.* 129 (3), 1085–1140.
- Allen, T., Donaldson, D., 2020. Persistence and Path Dependence in the Spatial Economy. Mimeo.
- Anderson, J., 1979. A theoretical foundation for the gravity equation. *Am. Econ. Rev.* 69 (1), 106–116.
- Armenter, R., Koren, M., Nagy, D., 2014. Bridges. Mimeo.
- Barjamovic, G., Chaney, T., Coşar, K., Hortaçsu, A., 2019. Trade, merchants, and the lost cities of the Bronze Age. *Q. J. Econ.* 134 (3), 1455–1503.
- Baum-Snow, N., Ferreira, F., 2015. Causal inference in urban and regional economics. In: Duranton, G., Henderson, V., Strange, W. (Eds.), *Handbook of Regional and Urban Economics*, vol. 5, pp. 3–68.
- Bolstad, P., 2005. *GIS Fundamentals: A First Text on Geographic Information Systems*. Eider Press.
- Bolt, J., Inklaar, R., de Jong, H., van Zanden, J., 2018. Rebasings 'Maddison': new income comparisons and the shape of long-run economic development. GGDC Res. Memorand. 174.

- Bos, F., 1992. The History of National Accounting. *MRPA Paper*. University Library of Munich.
- Brinkman, J., Lin, J., 2019. Freeway Revolts! the Quality of Life Effects of Highways. Mimeo.
- Brunet, M., 2010. Two new Mio-Pliocene Chadian hominids enlighten Charles Darwin's 1871 prediction. *Philos. Trans. R. Soc. B* 365 (1556), 3315–3321.
- Carson, C., 1975. The history of the United States national income and product accounts: the development of an analytical tool. *Rev. Income Wealth* 21, 153–181.
- Ciccone, A., Hall, R., 1996. Productivity and the density of economic activity. *Am. Econ. Rev.* 86 (1), 54–70.
- Combes, P., Gobillon, L., Zylberberg, Y., 2020. Urban Economics in a Historical Perspective: Recovering Data with Machine Learning. Mimeo.
- David, P., 1967. The growth of real product in the United States before 1840: new evidence, controlled conjectures. *J. Econ. Hist.* 27, 151–197.
- Delventhal, M., 2018. The Globe as a Network: Geography and the Origins of the World Income Distribution. Mimeo.
- Desmet, K., Nagy, D., Rossi-Hansberg, E., 2018. The geography of development. *J. Polit. Econ.* 126 (3), 903–983.
- Desmet, K., Rappaport, J., 2015. The settlement of the United States, 1800-2000: the long transition to Gibrat's Law. *J. Urban Econ.* forthcoming.
- Desmet, K., Rossi-Hansberg, E., 2014. Spatial development. *Am. Econ. Rev.* 104 (4), 1211–1243.
- Donaldson, D., 2018. Railroads of the Raj: Estimating the Impact of Transportation Infrastructure.
- Donaldson, D., Hornbeck, R., 2016. Railroads and American economic growth: a “market access” approach. *Q. J. Econ.* 131 (2), 799–858.
- Ducruet, C., Juhász, R., Nagy, D., Steinwender, C., 2020. All Aboard: the Effects of Port Development. Mimeo.
- Duranton, G., Morrow, P., Turner, M., 2014. Roads and trade: evidence from the US. *Rev. Econ. Stud.* 81 (2), 681–724.
- Duranton, G., Turner, M., 2011. The fundamental law of road congestion: evidence from US cities. *Am. Econ. Rev.* 101 (6), 2616–2652.
- Easterlin, R., 1960. Interregional differences in per capita income, population and total income, 1840–1950. In: *Trends in the American Economy in the Nineteenth Century*. Committee on Research in Income and Wealth. Princeton University Press.
- Eaton, J., Kortum, S., 2002. Technology, geography, and trade. *Econometrica* 70 (5), 1741–1779.
- Ebenstein, A., Fan, M., Greenstone, M., He, G., Yin, P., Zhou, M., 2015. Growth, pollution and life expectancy: China from 1991-2012. *Am. Econ. Rev.* 105 (5), 226–231.
- Eckert, F., Peters, M., 2018. Spatial Structural Change. Mimeo.
- Fajgelbaum, P., Redding, S., 2018. Trade, Structural Transformation and Development: Evidence from Argentina 1869-1914. *NBER Working Paper 20217*.
- Fajgelbaum, P., Schaal, E., 2020. Optimal transport networks in spatial equilibrium. *Econometrica* 88 (4), 1411–1452.
- Fariss, C., Crabtree, C., Anders, T., Jones, Z., Linder, F., Markowitz, J., 2017. Latent Estimation of GDP, GDP Per Capita, and Population from Historic and Contemporary Sources. Mimeo, <https://arxiv.org/pdf/1706.01099.pdf>.
- Fishbein, M., 1973. The Censuses of Manufactures 1810–1890. Reference Information Paper No. 50. National Archives and Records Service.
- Fogel, R., 1964. Railroads and American Economic Growth: Essays in Econometric History. Johns Hopkins University Press.
- Fujita, M., Thisse, J., 2002. Economics of Agglomeration. Cambridge University Press.
- Gallman, R., 1966. Gross national product in the United States, 1834–1909. In: NBER, Brady, D. (Ed.), *Output, Employment, and Productivity in the United States after 1800*, vol. 30. Columbia University Press. Studies in Income and Wealth.
- Galar, O., Özak, Ö., 2016. The agricultural origins of time preference. *Am. Econ. Rev.* 106 (10), 3064–3103.
- Hanlon, W., Hebllich, S., 2020. History and Urban Economics. Mimeo.
- Hebllich, S., Redding, S., Sturm, D., 2020a. The making of the modern metropolis: evidence from London. *Q. J. Econ.* 135 (4), 2059–2133.
- Hebllich, S., Trew, A., Zylberberg, Y., 2020b. East side story: historic pollution and neighborhood segregation. *J. Polit. Econ.* forthcoming.
- Herrendorf, B., Rogerson, R., Valentinyi, Á., 2014. Growth and structural transformation. *Handb. Econ. Growth* 2, 855–941 ch. 6.
- Kalweit, F., 1937. Die Baustellenwerte in Berlin: Strassen-Verzeichnis, Neuarbeitung Abgeschlossen Am 30. Juni 1936. Carl Heymanns.
- Klein Goldewijk, K., Beusen, A., Janssen, P., 2010. Long term dynamic modeling of global population and built-up area in a spatially explicit way: HYDE 3.1. Holocene 20, 565–573.
- Krugman, P., 1991. Increasing returns and economic geography. *J. Polit. Econ.* 99 (3), 483–499.
- Lee, S., Lin, J., 2018. Natural amenities, neighborhood dynamics, and persistence in the spatial distribution of income. *Rev. Econ. Stud.* 85 (1), 663–694.
- Lin, J., Rauch, F., 2020. What Future for History Dependence in Spatial Economics? Mimeo.
- Manson, S., Schroeder, J., van Riper, D., Ruggles, S., 2017. IPUMS National Historical Geographic Information System: Version 12.0. University of Minnesota, Minneapolis, <https://doi.org/10.18128/D050.V12.0>.
- Michaels, G., Rauch, F., Redding, S., 2019. Task specialization in U.S. cities from 1880-2000. *J. Eur. Econ. Assoc.* 17 (3), 754–798.
- Monte, F., Redding, S., Rossi-Hansberg, E., 2018. Commuting, migration and local employment elasticities. *Am. Econ. Rev.* 108 (12), 3855–3890.
- Nagy, D., 2020a. Hinterlands, City Formation and Growth: Evidence from the U.S. Westward Expansion. Mimeo.
- Nagy, D., 2020b. Trade and Urbanization: Evidence from Hungary. Mimeo.
- Pérez-Cervantes, F., 2014. Railroads and Economic Growth: A Trade Policy Approach. Banco de México Working Paper 2014-14.
- Peters, M., 2019. Market Size and Spatial Growth – Evidence from Germany's Post-war Population Expulsions. Mimeo.
- Ramankutty, N., Foley, J., Norman, J., McSweeney, K., 2002. The global distribution of cultivable lands: current patterns and sensitivity to possible climate change. *Global Ecol. Biogeogr.* 11 (5), 377–392.
- Redding, S., 2016. Goods trade, factor mobility and welfare. *J. Int. Econ.* 101, 148–167.
- Redding, S., Rossi-Hansberg, E., 2017. Quantitative spatial economics. *Ann. Rev. Econ.* 9, 21–58.
- Redding, S., Sturm, D., 2008. The costs of remoteness: evidence from German division and reunification. *Am. Econ. Rev.* 98 (5), 1766–1797.
- Redding, S., Sturm, D., 2016. Estimating Neighborhood Effects: Evidence from War-Time Destruction in London. Mimeo.
- Redding, S., Turner, M., 2015. Transportation costs and the spatial organization of economic activity. In: Duranton, G., Henderson, V., Strange, W. (Eds.), *Handbook of Regional and Urban Economics*, vol. 5, pp. 1339–1398.
- Santamaria, M., 2020. The Gains from Reshaping Infrastructure: Evidence from the Division of Germany. Mimeo.
- Shaw-Taylor, L., Wrigley, E., Davies, R., Kitson, P., Newton, G., Satchell, A., 2010. The Occupational Structure of England c.1710 to c.1871. Mimeo.
- Swisher, S., 2014. Reassessing Railroads and Growth: Accounting for Transport Network Endogeneity. Mimeo.
- Tompsett, A., 2020. Bridges: Economic Geography and the Impacts of Transport Infrastructure. Mimeo.
- Tóth, E., 2009. Duna-hídjaik. Közlekedésfejlesztési Koordinációs Központ.
- Trew, A., 2014. Spatial takeoff in the first industrial revolution. *Rev. Econ. Dynam.* 17, 707–725.
- Trew, A., 2020. Endogenous infrastructure development and spatial takeoff in the first industrial revolution. *Am. Econ. J. Macroecon.* 12 (2), 44–93.
- Ubelaker, D., 2008. Human Skeletal Remains: Excavation, Analysis, Interpretation. Transaction Pub.
- U.S. Census, 1841. Compendium of the Enumeration of the Inhabitants and Statistics of the United States. <http://usda.mannlib.cornell.edu/usda/AgCensusImages/1840/1840.pdf>.
- U.S. Census, 1852. The Seventh Census of the United States: 1850 - California. <https://www2.census.gov/library/publications/decennial/1850/1850a/1850a-47.pdf>.
- Weiss, T., 1992. U.S. labor force estimates and economic growth, 1800-1860. In: *American Economic Growth and Standards of Living before the Civil War*. University of Chicago Press, pp. 19–78.